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Children's Eating Attitudes Test (ChEAT): a validation study in Finnish children

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Abstract

Purpose To validate the Children's Eating Attitudes Test (ChEAT) in the Finnish population.

Materials and methods In total 339 children (age 10–15 years) from primary schools in Southern Finland were evaluated at two time points. They answered the ChEAT and SCOFF test questions, and had their weight, height and waist circumference measured. Retesting was performed 4–6 weeks later. Test–retest reliability was evaluated using intra-class correlation (ICC), and internal consistency was examined using Cronbach's alpha coefficient (C-alpha). ChEAT was cross-calibrated against SCOFF and background variables. Factor analysis was performed to examine the factor structure of ChEAT.

Results The 26-item ChEAT showed high internal consistency (C-alpha 0.79), however, a 24-item ChEAT showed even better internal consistency (C-alpha 0.84) and test–retest reliability (ICC 0.794). ChEAT scores demonstrated agreement with SCOFF scores ($p < 0.01$). The mean ChEAT score was higher in overweight children than normal weight ($p < 0.001$). Exploratory factor analysis yielded four factors (concerns about weight, limiting food intake, pressure to eat, and concerns about food), explaining 57.8% of the variance.

Conclusions ChEAT is a valid and reliable tool for measuring eating attitudes in Finnish children. The 24-item ChEAT showed higher reliability than the 26-item ChEAT.

Level of evidence Level 5, cross-sectional, descriptive study.

Keywords Adolescents · Children · Children's Eating Attitudes Test · ChEAT · Validation · Finland

Elisabete Weiderpass and Rejane Augusta de Oliveira Figueiredo equal contribution.

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Introduction

Increases in the prevalence of both obesity and disordered eating are a public health concern and might be associated [1–3]. Unhealthy eating attitudes and behaviors, such as dieting, vomiting and food avoidance, are considered disordered eating symptoms (DES) [4]. Eating disorder behaviors may eventually evolve into an eating disorder (ED) or a weight disorder (overweight and obesity), and therefore is of major public health relevance [5]. ED includes anorexia nervosa, bulimia nervosa, binge eating disorder, and other specified feeding and eating disorders [6].

Among 12- to 19-year-old adolescents, the prevalence of DES may range 9–17% [7, 8], while among 12–20 year olds the prevalence of ED has been estimated around 3–4% [9–11]. Although, in general, DES and ED are more common among females than males [10, 12], the female-to-male ratio increases during puberty [13].

DES in childhood and adolescence is likely to continue into young adulthood [12, 14] and is associated with being

overweight or obese, and having poor mental health and ED later in life [14–16]. Eating disturbance has been reported to start as early as the age of 9 years [17, 18]. For early recognition of DES, a reliable scale to identify DES in children is needed.

There are several instruments that can be used to assess DES in children, such as the SCOFF [19], the Eating Disorder Examination Questionnaire (EDE-Q) [20], the Eating Disorder Examination Questionnaire adapted for children (ChEDE-Q8) [21], the Eating Disorder Inventory for children (EDI-C) [22], the Kids' Eating Disorders Survey (KEDS) [23], and the Children's Eating Attitudes Test (ChEAT) [24]. ChEAT has been previously utilized in large cohort studies [25, 26], and it has been validated in Spain [27, 28], Portugal [29], Belgium [30], Japan [31], and the USA [24, 32, 33]. ChEAT has also been successfully used in Sweden [34]; however, to our knowledge, it has not yet been validated in Finland or in any of the other Nordic countries.

The aim of the present study is to validate ChEAT in a community sample of 10- to 15-year-old Finnish school children. This study is a part of the Finnish Health in Teens (Fin-HIT) cohort study.

Methods

Participants

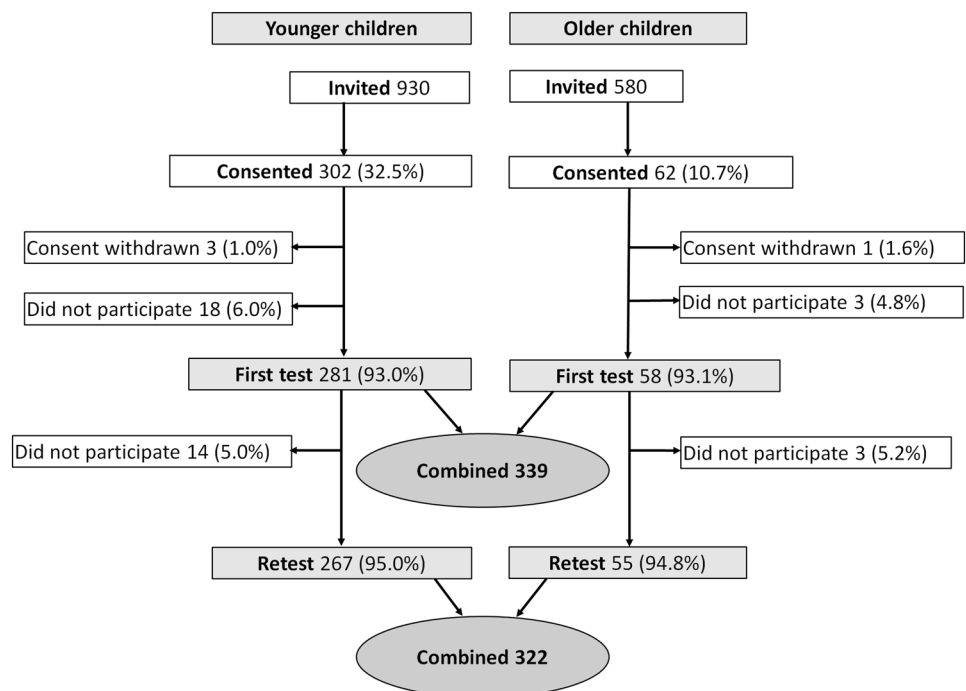
The present study, conducted by Folkhälsan Research Center in Helsinki, was initiated to validate ChEAT among Finnish

children. ChEAT was used in the Fin-HIT cohort to evaluate disordered eating symptoms [35], and it was chosen as it had been successfully validated in several countries, as described above, and utilized in several previous studies (e.g., [17, 34, 36]). The Fin-HIT is a prospective cohort with approximately 11,000 9- to 14-year-old participants across Finland. The proposed validation sub-study consisted of 339 children at the same age as children in the Fin-HIT cohort at baseline. Figure 1 shows the participant selection flowchart. Participants were recruited from 12 primary schools in the Helsinki metropolitan area, and the data collection was conducted between October 2017 and March 2018. After receiving recruitment permission from municipal education administration and the principal of each school, fieldworkers introduced the study protocol and distributed invitations to children in suitable classes. To participate, children and one parent per each child provided a written informed consent. The Coordinating Ethics Committee of the Hospital District of Helsinki and Uusimaa has approved the study protocol.

Data collection

The data collection was conducted at two time points during normal school days: an initial visit and a retest visit 4–6 weeks later. Participants answered a questionnaire consisting of ChEAT, a question related to overeating symptoms (frequency of having eaten too much within a short time) [37], SCOFF and a short food frequency questionnaire. Instructions to fill in the questionnaire were explained simultaneously to all participants, and possible questions arising

Fig. 1 Flowchart of study



afterward were answered individually. After completing the questionnaire, at the first visit only, participants' weight, height and waist circumference were measured by trained fieldworkers. Due to absence from school, some participants filled in the questionnaire and took their measurements at home. We have previously shown the validity of home-based measures of weight, height, and waist circumference, and these were found to be reliable [38]. Of the 339 participants who attended the first data collection, 6 participated at home, and of the 322 retested participants, 16 participated at home.

Study sample size

Sample size calculations were based on ChEAT scores obtained from children in the Fin-HIT cohort (mean 6.1, standard deviation 4.7) [25]. Adopting a significance level of 5% and a power of 80%, the minimum required sample size calculated was 214 children, allowing us to compare the results between the two time points (test and retest) and to test the internal consistency of the scale. We recruited children in 5th to 8th grades, but since the response rate was low among children in 7th and 8th grades (hereafter referred to as older children), we chose to carry out the study with children in 5th and 6th grades (hereafter referred to as younger children). The adequate sample size was met with younger children. The 58 older children were included as a complementary sample.

Measures

ChEAT

ChEAT is a self-administered questionnaire for assessing eating attitudes and behaviors in children, based on the Eating Attitudes Test (EAT-26) used in adults [39]. Since some wordings of EAT-26 items were deemed incomprehensible to children, a simplified version was developed and validated [24]. ChEAT consists of 26 items scored on a 6-point Likert scale with answer options 'never', 'rarely', 'sometimes', 'often', 'usually', and 'always'. To calculate the scale's scores, the categories 'never', 'rarely', and 'sometimes' were scored as zero, 'often' as one, 'usually' as two, and 'always' as three. Item 19 was reverse-scored, as suggested by Maloney [24]. The maximum score is 78 points. The higher the score, the higher the possibility for DES. The original ChEAT had three subscales: "Dieting", "Bulimia and Food Preoccupation" and "Oral Control", respectively, with 13, 6, and 7 items [39]. ChEAT was created and validated in English. The instrument used in Fin-HIT cohort and in the present study was evaluated by the research group and the final version was translated into Finnish by linguistic professionals. The scale adaptation for Finnish-speaking children was performed at the beginning of the Fin-HIT study.

SCOFF

The questionnaire also included SCOFF, a screening tool for detecting DES, consisting of five items that are dichotomously scored (yes = 1, no = 0). Two or more "yes" responses indicate DES [19]. The acronym SCOFF is composed of the following questions: (1) do you make yourself Sick because you feel uncomfortably full?; (2) do you worry you have lost Control over how much you eat?; (3) have you recently lost more than 6 kg (one stone) in a 3-month period?; (4) do you believe yourself to be Fat when others say you are too thin?; and (5) would you say that Food dominates your life? SCOFF was used to assess the concurrent validity since it was validated and used in several Finnish studies [40, 41].

Anthropometric measurements

Weight, height and waist circumference were measured by field workers in a standardized way. Weight was measured to the nearest 0.01 kg with a portable digital scale (CAS model PB), and height to the nearest 0.1 cm with a portable stadiometer (Seca model 217). Digital scale was calibrated daily before each series of measurements. Before measurements, the participants removed their shoes, heavy sweaters, and any objects from their pockets. Fieldworkers reported the clothing participants were wearing, and the weight of the clothes was subtracted from the measured weight in a standardized way. Waist circumference was measured midway between the hip bones and ribs to the nearest 0.1 cm with a measuring tape calibrated against a measure approximately once a week. Waist measurement was repeated twice, and the recorded waist circumference was calculated as their mean. In cases with more than 1 cm difference between the two measurements, a third measurement was performed and the mean was calculated based on all three measurements.

Body mass index (BMI) was calculated as weight (kg)/height (m)², and the children were classified as underweight, normal weight, overweight, or obese according to age- and sex-specific cut-offs suggested by the International Obesity Task Force (IOTF) [42]. Information on age and gender was obtained from the consent form and the questionnaire.

Statistical analysis

All associations between categorical variables were assessed using Chi-square tests. Comparison of quantitative variables with dichotomous variables was carried out using Student's *t* test. ChEAT was evaluated for 300 children who answered all ChEAT items. Comparisons between ChEAT score and BMI categories were performed by generalized linear models using Poisson distribution. The internal consistency of all scale items was assessed using Cronbach's alpha coefficient

Table 1 Characteristics of all children and separated by age groups

| | All children | | Age group | | | | <i>p</i> value |
|--------------------------|--------------|------|------------------|------|----------------|-------|----------------|
| | | | Younger children | | Older children | | |
| | <i>n</i> | % | <i>n</i> | % | <i>n</i> | % | |
| <hr/> | | | | | | | |
| Gender | | | | | | | |
| Girls | 185 | 54.6 | 155 | 55.2 | 30 | 51.7 | 0.632 |
| Boys | 154 | 45.4 | 126 | 44.8 | 28 | 48.3 | |
| BMI | | | | | | | |
| Underweight | 41 | 12.1 | 36 | 12.8 | 5 | 8.6 | 0.664 |
| Normal weight | 246 | 72.6 | 201 | 71.5 | 45 | 77.6 | |
| Overweight | 49 | 14.5 | 41 | 14.6 | 8 | 13.8 | |
| Obese | 3 | 0.9 | 3 | 1.1 | 0 | 0.0 | |
| School size | | | | | | | |
| < 300 students | 91 | 26.8 | 91 | 32.4 | 0 | 0.0 | <0.001 |
| > 300 students | 248 | 73.2 | 190 | 67.6 | 58 | 100.0 | |
| DES (SCOFF ≥ 2) | | | | | | | |
| No DES | 284 | 89.3 | 232 | 88.2 | 52 | 94.5 | 0.167 |
| DES | 34 | 10.7 | 31 | 1.8 | 3 | 5.5 | |
| <hr/> | | | | | | | |
| | Mean | SD | Mean | SD | Mean | SD | |
| <hr/> | | | | | | | |
| ChEAT score (26 items) | 2.55 | 4.22 | 2.66 | 4.44 | 2.05 | 2.84 | 0.319 |
| Age (years) | 11.9 | 0.9 | 11.6 | 0.6 | 13.3 | 0.7 | <0.001 |
| Waist circumference (cm) | 65.4 | 7.4 | 64.6 | 7.2 | 69.3 | 7.1 | <0.001 |
| Weight (kg) | 44.6 | 10.3 | 42.8 | 9.1 | 53.0 | 11.6 | <0.001 |

BMI body mass index, *DES* disordered eating symptoms, *SD* standard deviation

Table 2 Results of the test and retest analysis for all children and separated by age groups

| | ChEAT score (test) | | ChEAT score (retest) | | Paired <i>t</i> test (<i>p</i> value) | Pearson correlation | ICC |
|------------------|--------------------|------|----------------------|------|--|---------------------|-------|
| | Mean | SD | Mean | SD | | | |
| 26-item ChEAT | | | | | | | |
| All children | 2.55 | 4.22 | 2.40 | 3.91 | 0.331 | 0.771 | 0.769 |
| Younger children | 2.66 | 4.44 | 2.45 | 4.10 | 0.266 | 0.780 | 0.779 |
| Older children | 2.05 | 2.84 | 2.16 | 2.83 | 0.774 | 0.680 | 0.670 |
| 24-item ChEAT | | | | | | | |
| All children | 2.03 | 4.13 | 1.80 | 3.81 | 0.119 | 0.796 | 0.794 |
| Younger children | 2.12 | 4.37 | 1.86 | 3.99 | 0.129 | 0.797 | 0.795 |
| Older children | 1.57 | 2.73 | 1.53 | 2.83 | 0.717 | 0.784 | 0.784 |

ICC intra-class correlation, *SD* standard deviation

(C-alpha), evaluating the correlation of each item with all the other items together.

To evaluate the factorial structure of all items and to identify different subscales in ChEAT, we used exploratory factor analysis. A model's goodness-of-fit was evaluated by the Kaiser–Meyer–Olkin (KMO) and Bartlett's sphericity tests, considering values over 0.70 and $p < 0.05$, respectively, as acceptable [43]. Once the factors were identified, confirmatory factor analysis (CFA) was used to check the scale's factor structures. Estimations for CFAs were carried out using

a robust unweighted least squares method. The goodness-of-fit of these two models was assessed using the following measures: Chi-square test, comparative fit index (CFI), non-normed fit index (NNFI), root mean square error (RMSEA), standardized mean square residual (SRMR), goodness-of-fit index (GFI), and adjusted goodness-of-fit index (AGFI) [44].

Test–retest results were evaluated using paired *t* tests to compare mean scores, Pearson correlation and intra-class correlation (ICC) to evaluate reliability and homogeneity of results between the two time points.

Table 3 Comparison of ChEAT score by SCOFF score, binge eating symptom, gender, and body mass index (BMI) categories for all children and separated by age groups

| | All | | | <i>p</i> value | Younger | | | <i>p</i> value | Older | | | <i>p</i> value |
|--|----------|------|------|----------------------|----------|------|------|----------------------|----------|------|------|----------------|
| | <i>n</i> | Mean | SD | | <i>n</i> | Mean | SD | | <i>n</i> | Mean | SD | |
| 26-item ChEAT | | | | | | | | | | | | |
| SCOFF | | | | | | | | | | | | |
| No DES | 246 | 1.85 | 2.58 | < 0.001 | 198 | 1.87 | 2.58 | 0.010 | 48 | 1.79 | 2.58 | 0.005 |
| DES | 31 | 6.94 | 9.29 | | 28 | 6.96 | 9.67 | | 3 | 6.67 | 5.51 | |
| Frequency of having eaten too much in short time | | | | | | | | | | | | |
| Never | 212 | 1.94 | 2.94 | 0.013 | 174 | 1.90 | 2.93 | 0.009 | 38 | 2.13 | 3.00 | 0.769 |
| Sometimes in a month | 79 | 3.68 | 5.89 | | 64 | 4.11 | 6.34 | | 15 | 1.87 | 2.80 | |
| Gender | | | | | | | | | | | | |
| Girls | 163 | 2.88 | 5.04 | 0.061 | 134 | 3.06 | 5.39 | 0.051 | 29 | 2.07 | 2.83 | 0.971 |
| Boys | 135 | 1.99 | 3.05 | | 110 | 1.98 | 3.07 | | 25 | 2.04 | 3.02 | |
| BMI | | | | | | | | | | | | |
| Underweight | 36 | 2.00 | 2.06 | < 0.001 ^a | 32 | 2.16 | 2.08 | < 0.001 ^a | 4 | 0.75 | 1.50 | 0.150 |
| Normal | 218 | 2.31 | 3.41 | | 176 | 2.36 | 3.56 | | 42 | 2.10 | 2.68 | |
| Overweight/obese | 44 | 3.70 | 7.88 | | 36 | 3.97 | 8.49 | | 8 | 2.50 | 4.38 | |
| 24-item ChEAT | | | | | | | | | | | | |
| SCOFF | | | | | | | | | | | | |
| No DES | 284 | 1.47 | 2.50 | 0.004 | 232 | 1.50 | 2.54 | 0.009 | 52 | 1.33 | 2.34 | 0.001 |
| DES | 34 | 6.29 | 9.14 | | 31 | 6.26 | 9.48 | | 3 | 6.67 | 5.51 | |
| Frequency of having eaten too much in short time | | | | | | | | | | | | |
| Never | 243 | 1.52 | 2.90 | 0.009 | 202 | 1.52 | 2.92 | 0.008 | 41 | 1.54 | 2.81 | 0.919 |
| Sometimes in a month | 88 | 3.22 | 5.69 | | 72 | 3.57 | 6.12 | | 16 | 1.63 | 2.68 | |
| Gender | | | | | | | | | | | | |
| Girls | 185 | 2.27 | 4.83 | 0.235 | 155 | 2.43 | 5.12 | 0.186 | 30 | 1.43 | 2.79 | 0.999 |
| Boys | 154 | 1.73 | 3.10 | | 126 | 1.74 | 3.19 | | 28 | 1.71 | 2.71 | |
| BMI | | | | | | | | | | | | |
| Underweight | 41 | 1.32 | 2.13 | < 0.001 ^b | 36 | 1.50 | 2.21 | < 0.001 ^a | 5 | 0.00 | 0.00 | 0.069 |
| Normal | 246 | 1.80 | 3.22 | | 201 | 1.85 | 3.36 | | 45 | 1.58 | 2.48 | |
| Overweight/obese ^a | 52 | 3.67 | 7.52 | | 44 | 3.89 | 7.98 | | 8 | 2.50 | 4.38 | |

SD standard deviation, DES disordered eating symptoms, BMI body mass index

^aMultiple comparison between overweight and other groups $p < 0.001$

^bMultiple comparison between all BMI categories $p < 0.05$

We used the Lavaan package in R for CFA and SPSS statistical software (version 24.0) for all other statistical analyses. We adopted a 5% statistical significance level for all tests.

Results

Descriptive

Of the 339 participants, 82.9% ($n = 281$) were younger and 17.1% ($n = 58$) older children. The overall mean age of participants was 11.9 (SD ± 0.9) years: 11.6 (± 0.6) years for younger and 13.3 (± 0.7) years for older children. Overall, 54.6% ($n = 185$) were girls and 45.4% ($n = 154$) were boys.

Table 1 shows the main characteristics of all participants, and comparisons between younger and older children. Gender, BMI and presence of disordered eating did not differ between younger and older children.

Internal consistency

Cronbach's alpha coefficient (C-alpha) showed high internal consistency (C-alpha 0.79) for the 26-item ChEAT, however, the reverse-scored item 19 (I can show self-control around food) and the item 25 (I enjoy trying new rich foods) had a negative correlation with other items. A negative correlation indicates a reverse direction of these items related to the other items, suggesting that they should not be included in the final scale (called the 24-item ChEAT).

After exclusion of the two items, the internal consistency increased (C-alpha 0.84). The internal consistency was also evaluated separately for younger and older children, and there was improvement in the C-alpha after exclusion in both groups: increasing from 0.81 to 0.84 in younger children, and from 0.64 to 0.73 in older children.

Test and retest

Table 2 shows results for test and retest analysis. We found no statistical difference between test and retest in the mean total ChEAT scores, either with the 26-item ChEAT ($p=0.331$) or the 24-item ChEAT ($p=0.119$). The results remained similar when stratified by age group. When evaluating the homogeneity between the two time points using ICC, the overall homogeneity was higher with the 24-item ChEAT (ICC 0.794) than with the 26-item ChEAT (ICC 0.769). Similar results were obtained when separated by age group (Table 2).

Cross-calibrating ChEAT against SCOFF and other variables

We compared the mean score of ChEAT with the mean score of SCOFF and with the presence of the overeating symptom (frequency of having eaten too much in a short time) using the t test (Table 3). We found that the mean ChEAT score is higher in children with DES based on SCOFF than in children without DES ($p<0.01$). Similar results were observed among the different age groups. Children with overeating at least once a month had higher ChEAT scores than children without overeating ($p=0.013$). We obtained the same results among younger children ($p=0.009$), but not with older children ($p=0.769$).

Mean ChEAT scores were also compared between genders and BMI groups. The mean ChEAT score did not differ between genders ($p=0.061$), but did differ between BMI categories ($p<0.001$): higher scores were related with higher BMI. Similar results were obtained when using the 24-item ChEAT score (Table 3).

Exploratory and confirmatory factor analysis

After excluding the two items from the scale, we performed a factor analysis to evaluate the factorial structure. Factor analysis identified four subscales (i.e., four factors; Table 4) with high adaptability to the original data ($KMO=0.775$ and $p<0.001$ for Bartlett's sphericity test), and these explained 57.8% of the data variability.

We used CFA to verify the scale structure with three different approaches: (1) a model with the original structure; (2) a model with the original structure excluding

Table 4 Loads for new 24-item ChEAT factors after exploratory factor analysis

| | Factor 1 | Factor 2 | Factor 3 | Factor 4 |
|---------|----------|----------|----------|----------|
| Item 1 | 0.494 | 0.073 | 0.191 | 0.036 |
| Item 6 | 0.438 | -0.075 | 0.347 | 0.202 |
| Item 7 | 0.614 | 0.343 | 0.064 | 0.160 |
| Item 10 | 0.814 | -0.025 | 0.182 | -0.063 |
| Item 11 | 0.630 | 0.329 | 0.117 | -0.061 |
| Item 12 | 0.534 | 0.148 | 0.242 | 0.165 |
| Item 13 | 0.307 | -0.005 | 0.182 | 0.091 |
| Item 14 | 0.786 | -0.054 | -0.032 | -0.089 |
| Item 17 | 0.600 | 0.114 | -0.106 | 0.426 |
| Item 18 | 0.421 | 0.089 | 0.017 | 0.165 |
| Item 22 | 0.427 | -0.052 | -0.020 | 0.231 |
| Item 23 | 0.542 | 0.187 | -0.055 | 0.137 |
| Item 24 | 0.316 | 0.195 | -0.046 | 0.151 |
| Item 2 | 0.141 | 0.586 | 0.424 | 0.100 |
| Item 4 | 0.418 | 0.660 | -0.043 | 0.270 |
| Item 5 | 0.013 | 0.607 | 0.069 | -0.120 |
| Item 9 | 0.034 | 0.642 | -0.081 | 0.365 |
| Item 26 | 0.048 | 0.894 | 0.081 | -0.077 |
| Item 8 | 0.001 | 0.020 | 0.526 | -0.089 |
| Item 15 | 0.178 | 0.276 | 0.451 | -0.086 |
| Item 20 | 0.380 | 0.364 | 0.455 | 0.149 |
| Item 21 | -0.023 | -0.051 | 0.748 | 0.258 |
| Item 3 | 0.092 | 0.061 | 0.024 | 0.715 |
| Item 16 | 0.159 | 0.011 | 0.096 | 0.332 |

items 19 and 25; and (3) a model with a new structure with four factors as suggested by the exploratory factor analysis (Table 5). In all children, model 1 showed the poorest goodness-of-fit compared with models 2 and 3 for all measures. Both model 2 and 3 showed high goodness-of-fit in all children, according to the following measures: CFI, NNFI, RMSEA, Chi-square, GFI and AGFI [44]. Similar results were obtained in the subgroup of younger children. The sample size was not large enough to evaluate models separately in the older children. After checking the goodness-of-fit of model 3, we evaluated the standard factor loads estimated by the CFA (Table 6), interpreted the results and named the four factors accordingly: subscale 1 as "Concerns about weight"; subscale 2 as "Limiting food intake"; subscale 3 as "Pressure to eat"; and subscale 4 as "Concerns about food". To cross-validate our results, we carried out the CFA with 11,407 children from Fin-HIT cohort, and found similar results: even though the results for the 26-item scale (model 1) were good, the results for the 24-item scale were even better (model 2 and 3) (Table 5).

Table 5 Results of confirmatory factorial analysis for three different models evaluated in the study for all children and younger children

| | Model 1 (26 items—original) | Model 2 (24 items—original) | Model 3 (24 items—4 new factors) | Criteria [44] |
|---|--------------------------------|--------------------------------|--|-------------------------|
| All children | | | | |
| CFI | 0.872 | 0.959 | 0.955 | CFI \geq 0.90 |
| NNFI | 0.859 | 0.953 | 0.948 | NNFI \geq 0.90 |
| RMSEA | 0.028 | 0.016 | 0.017 | RMSEA $<$ 0.08 |
| SRMR | 0.101 | 0.102 | 0.091 | SRMR $<$ 0.08 |
| Chi-square | 0.003 | 0.167 | 0.144 | <i>p</i> value $>$ 0.05 |
| GFI | 0.976 | 0.974 | 0.971 | GFI \geq 0.95 |
| AGFI | 0.970 | 0.966 | 0.963 | AGFI \geq 0.90 |
| Younger children | | | | |
| CFI | 0.842 | 0.941 | 0.959 | CFI \geq 0.90 |
| NNFI | 0.826 | 0.932 | 0.953 | NNFI \geq 0.90 |
| RMSEA | 0.032 | 0.021 | 0.017 | RMSEA $<$ 0.08 |
| SRMR | 0.112 | 0.113 | 0.102 | SRMR $<$ 0.08 |
| Chi-square | 0.001 | 0.110 | 0.194 | <i>p</i> value $>$ 0.05 |
| GFI | 0.970 | 0.971 | 0.966 | GFI \geq 0.95 |
| AGFI | 0.962 | 0.960 | 0.957 | AGFI \geq 0.90 |
| Cross-validation analysis—using Fin-HIT data (<i>n</i> = 11,407) | | | | |
| CFI | 0.950 | 0.969 | 0.991 | CFI \geq 0.90 |
| NNFI | 0.891 | 0.965 | 0.944 | NNFI \geq 0.90 |
| RMSEA | 0.041 | 0.043 | 0.038 | RMSEA $<$ 0.08 |
| SRMR | 0.054 | 0.056 | 0.057 | SRMR $<$ 0.08 |
| Chi-square | 0.098 | 0.076 | 0.072 | <i>p</i> value $>$ 0.05 |
| GFI | 0.990 | 0.980 | 0.986 | GFI \geq 0.95 |
| AGFI | 0.989 | 0.976 | 0.982 | AGFI \geq 0.90 |

CFI comparative fit index, NNFI non-normed fit index, RMSEA root mean square error, SRMR standardized mean square residual, GFI goodness-of-fit index, AGFI adjusted goodness-of-fit index

Discussion

The present study validated the ChEAT scale in a population of Finnish children and proposed a new 24-item ChEAT. We identified new subscales to evaluate different aspects of DES in children. Validation was conducted with different approaches, all showing good consistency and reliability.

We evaluated the internal consistency of the original 26-item scale and observed two items (19 ‘I can show self-control around food’ and 25 ‘I enjoy trying new rich foods’) having an inverse correlation with all other items, indicating a semantic inversion of the statements, thus suggesting that these items were not clear for all children. The inverse correlation of the reverse-scored item 19 could be due to response bias, since the participants may have had a tendency to reply to this item in the same direction as to the other items. By excluding these two items, the reliability of the scale improved. The CFA reinforced these findings, since a better goodness-of-fit was observed with the 24-item scale. The exclusion of one or both of these items

have been suggested by others too [24, 30, 31, 33], resulting in improved consistency.

We cross-calibrated ChEAT with other variables that have been associated with DES in previous studies. The SCOFF is a widely used, validated screening tool in clinical settings in Finland [41, 45, 46]. We observed that the mean ChEAT score was higher among those who scored higher in SCOFF. The ChEAT also discriminated between children with overeating and without overeating. Previous studies show that overweight children are more likely to have DES [47, 48], and our result of overweight children having higher ChEAT scores than normal weight children was in line with these. We also compared ChEAT score by gender but did not find any difference, which is consistent with results from some earlier studies [28, 49]. However, DES is shown to be more frequent in girls than boys in some studies [50, 51], but the gender difference is proposed to emerge around the age of 13 [52–54].

The mean ChEAT scores in our sample were lower than observed in previous studies [31, 33, 55]. The majority of

Table 6 Confirmatory factorial analysis results for 24-item ChEAT

| Item | Standardized estimation |
|--|-------------------------|
| Subscale 1 “concerns about weight” | |
| 1—I am scared about being overweight | 0.451 |
| 6—I am aware of the energy (calorie) content in foods that I eat | 0.538 |
| 7—I try to stay away from foods such as breads, potatoes, and rice | 0.536 |
| 10—I feel very guilty after eating | 0.703 |
| 11—I think a lot about wanting to be thinner | 0.615 |
| 12—I think about burning up energy (calories) when I exercise | 0.664 |
| 13—Other people think I am too thin | 0.310 |
| 14—I think a lot about having fat on my body | 0.624 |
| 17—I eat diet foods | 0.558 |
| 18—I think that food controls my life | 0.398 |
| 22—I feel uncomfortable after eating sweets | 0.435 |
| 23—I have been dieting | 0.517 |
| 24—I like my stomach to be empty | 0.348 |
| Subscale 2 “limiting food intake” | |
| 2—I stay away from eating when I am hungry | 0.800 |
| 4—I have gone on eating binges where I feel that I might not be able to stop | 0.690 |
| 5—I cut my food into small pieces | 0.389 |
| 9—I vomit after I have eaten | 0.483 |
| 26—I feel I have to vomit after a meal | 0.604 |
| Subscale 3 “pressure to eat” | |
| 8—I feel that others would like me to eat more | 0.237 |
| 15—I take longer than others to eat my meals | 0.508 |
| 20—I feel that others pressure me to eat | 0.824 |
| 21—I give too much time and thought to food | 0.395 |
| Subscale 4 “concerns about food” | |
| 3—I think about food a lot of the time | 0.428 |
| 16—I stay away from foods with sugar in them | 0.503 |

our participants lived in a small or medium-sized city or in the rural areas of a large city, which could partly explain our findings since mean ChEAT scores have been observed to be higher in large cities than in small and medium-sized cities, possibly due to susceptibility to more social pressure and to emphasized media influence in larger cities [31]. It has been suggested that socioeconomic status affects the prevalence of DES [56, 57], which might explain low mean scores in our study. However, information on parental SES was not available here.

Our exploratory factor analysis yielded four factors or subscales that we named “Concerns about weight”, “Limiting food intake”, “Pressure to eat”, and “Concerns about food” that describe relevant attitudes and behaviors. The factor structure was confirmed in the CFA. In previous studies, the typical number of factors has been four or five, which is in line with our results [27, 31, 55, 58–60]. In some previous studies [27, 31, 61], factor analysis revealed a factor representing behaviors or thoughts related to purging. In our analysis, purging (‘I vomit after I have eaten’ and ‘I have the

urge to vomit after a meal’) did not stand out as an independent factor, but rather was associated with behaviors related to controlling the amount of food eaten.

Typically, validation studies are performed without proper sample size calculations [62], but this study was performed with an adequate sample size. We used several statistical approaches to validate ChEAT. Complementarily, we cross-calibrated ChEAT against several background variables that included SCOFF, overeating symptom, BMI, and gender. Furthermore, a cross-validation analysis in 11,407 Fin-HIT participants confirmed the results of the CFA. Still, this study had some limitations. The low participation rate may cause some bias. The participants are likely to be more health-oriented than the non-participants [63, 64], which could be the reason for low mean ChEAT scores in general. The low participation rate might also be due to a lack of a systematic reminder to parents to participate. Teachers were asked to send a reminder to parents through a messaging system, but it was not ensured whether they actually sent them. Since we did not have ED diagnoses available, we

were not able to evaluate a ChEAT cut-off for identifying DES. The number of studies investigating a suitable cut-off for children is limited [27, 28, 31]. Cut-offs varying from 10 to 20 have been tested in age groups between 9 and 17 [27, 28, 31], but there is yet no consensus on an optimal cut-off. The Fin-HIT cohort will include data on ED diagnoses from national healthcare registers, thus it will be possible to determine a suitable cut-off for children in the future. Moreover, the restricted range of the ChEAT scoring could be a possible limitation in our study. However, the restricted range reduces correlation coefficient values [65], which in this study is not a problem as we yielded good correlation coefficients even with the restricted range.

To conclude, we assessed the consistency and reliability of the ChEAT using an adequate sample size among Finnish children. ChEAT is a valid and reliable tool to measure eating attitudes and behaviors in Finnish children in 5th and 6th grades (10–13 years). We proposed a 24-item ChEAT to be used in future studies among Finnish children.

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Compliance with ethical standards

Conflict of interest The authors declare that they have no conflict of interest.

Ethical approval All procedures performed in studies involving human participants were in accordance with the ethical standards of the institutional and/or national research committee and with the 1964 Helsinki declaration and its later amendments or comparable ethical standards. The Coordinating Ethics Committee of the Hospital District of Helsinki and Uusimaa has approved the study protocol (169/13/03/00/10).

Informed consent Informed consent was obtained from all individual participants included in the study.

Permission The ChEAT has been reproduced with permission from Maloney et al. (1988). Reliability Testing of a Children's Version of the Eating Attitudes Test. *Journal of the American Academy of Child and Adolescent Psychiatry*.

Data availability The datasets generated and/or analyzed during the current study are available from the corresponding author on reasonable request.

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